

# Factorial Moments of Continuous Order and Multifractal Analysis in 400 GeV/c pp Collisions

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The method of factorial moments  $F_q$  of continuous order suggested by Hwa has been tested. It was found that using this method to analyze the experimental data will not produce satisfactory results. Some improvements were made for Hwa's method, making it suitable for multifractal analysis of experimental data. The analytic results for the experimental pseudorapidity distributions of charged particles produced in 400 GeV/c pp collisions indicated that the method of the factorial moments of continuous order is feasible. There is possibly multifractal behavior in the process of multiplicity production in pp collisions at 400 GeV/c.

**Key words:** factorial moments of continuous order, negative binomial distribution, maximum likelihood function method, multifractal behavior.

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## 1. INTRODUCTION

In order to investigate the intermittent phenomenon of multiparticle productions in high energy collisions, one may calculate the moments of the multiplicity distribution and study their dependence on the size of phase space. The multiplicity distribution originates from the dynamic and statistical fluctuations. In high energy hadron collisions, the multiplicity is usually small, especially when the phase space becomes small, and the statistical fluctuation will be a dominant factor compared to the dynamic one. Bialas and Peschanski suggested investigating the intermittent phenomenon using the scaled factorial moment  $F_q$  [1] which was defined as

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$$F_q = \frac{\langle n(n-1)\cdots(n-q+1) \rangle}{\langle n \rangle^q}, \quad (1)$$

which gives a nonbiased estimation of the moments of dynamic fluctuation. It greatly improved the study of intermittency. However, the defect of the factorial moment is that the order of moment should be an integer greater than or equal to 2. Consequently, one cannot obtain two important fractal dimensions  $D_0$  and  $D_1$  and the dip signal of the pseudorapidity distributions. It was later suggested to use the  $G_q$  moments and various modified ones, which have arbitrary  $q$  values and overcome the defects of  $F_q$ . However, the  $G_q$  moments have an inherent defect, i.e., they did not eliminate the influence of statistical background. Subtraction of the statistical component has to be done by hand [2]. Because of this, it is necessary to make the order of  $F_q$  moments continuously varying in order to have both merits of  $F_q$  and  $G_q$  and to overcome both defects. Hwa recently proposed a new method [3] to obtain  $F_q$  moment of continuous order. However, it is unsatisfactory for use in experimental data analysis. We have made some modifications for this method and made it feasible for multifractal analysis of experimental data.

## 2. HWA'S METHOD OF FACTORIAL MOMENT OF CONTINUOUS ORDER

The multiplicity distribution  $P_n$  can be represented as

$$P_n = S \times D, \quad (2)$$

where  $S$  represents the statistical fluctuation which was described by Poisson distribution, and  $D$  represents the dynamic component. Further, it can be written as

$$P_n = \int_0^\infty dt \frac{t^n}{n!} e^{-t} D(t). \quad (3)$$

The scaled factorial moment  $F_q$  can be written as

$$F_q = f_q / f_1^q, \quad (4)$$

where

$$f_q = \sum_{n=q}^{\infty} \frac{n!}{(n-q)!} P_n. \quad (5)$$

Because  $q$  is a positive integer, substituting Eq. (3) into Eq. (5) and performing the summation over  $n$ , one can obtain

$$f_q = \int_0^\infty dt t^q D(t). \quad (6)$$

It is the  $q$ -th moment of the dynamic fluctuation  $D(t)$ . However, for continuous order  $q$ , Eq. (5) cannot be applied to obtain Eq. (6) of the  $q$ -th moment  $f_q$  of the dynamic fluctuation. So, other methods are needed. It is demanded in Ref. [3] that Eq. (6) is suitable for all  $q$ . In order to obtain the expression of  $D(t)$ , the multiplicity distribution  $P_n$  is expanded using negative binomial distribution (NBD), i.e.,

$$P_n = \sum_{j=0}^N a_j P_n^{\text{NB}}(k_j, x_j), \quad (7)$$

where  $n = 0, 1, \dots, N$  ( $N$  is the maximum multiplicity), and

$$P_n^{\text{NB}}(k_j, x_j) = \frac{\Gamma(n + k_j)}{\Gamma(n + 1) \Gamma(k_j)} \left(\frac{k_j}{k_j + x_j}\right)^{k_j} \left(\frac{x_j}{k_j + x_j}\right)^n. \tag{8}$$

Using Eq. (3) and the following relation

$$P_n^{\text{NB}}(k_j, x_j) = \int_0^\infty dt \frac{t^n}{n!} e^{-t} D^{\text{NB}}(t, j), \tag{9}$$

where

$$D^{\text{NB}}(t, j) = \left(\frac{k_j}{x_j}\right)^{k_j} \frac{t^{k_j-1}}{\Gamma(k_j)} e^{-(k_j + t/x_j)}, \tag{10}$$

one gets

$$D(t) = \sum_{j=0}^N a_j D^{\text{NB}}(t, j). \tag{11}$$

Substituting (11) to Eq. (6) and integrating over  $t$ , one finds

$$f(q) = \sum_{j=0}^N a_j f^{\text{NB}}(q, j), \tag{12}$$

where

$$f^{\text{NB}}(q, j) = \left(\frac{x_j}{k_j}\right)^q \frac{\Gamma(q + k_j)}{\Gamma(k_j)}. \tag{13}$$

Then one obtains the relation similar to Eq. (4)

$$F(q) = f(q) / f(1)^q. \tag{14}$$

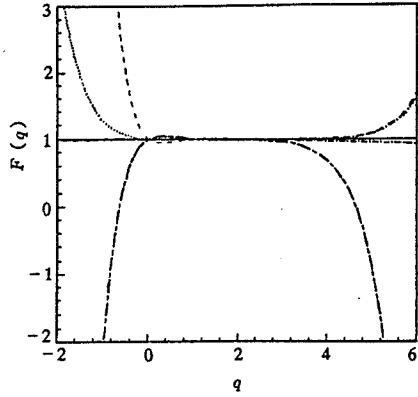
To determine the value of  $a_j$ , one should assign  $N + 1$  pairs of  $x_j, k_j$ , then calculate  $P_n^{\text{NB}}(k_j, x_j)$  according to Eq. (8), and get the values of  $a_j$  by solving the  $N + 1$  linear algebraic Eq. (7). Therefore, the factorial moments of continuous order  $F(q)$  can be obtained by Eqs. (12) and (14).

### 3. TEST OF HWA'S METHOD

To have a test of Hwa's method, we apply Hwa's method to the multiplicity distribution with Poissonian distribution

$$P_n = e^{-\langle n \rangle} \langle n \rangle^n / n!, \tag{15}$$

$P_n$  is calculated analytically according to the above equation with  $\langle n \rangle = 1, n = 0, 1, \dots, N; N = 10$ . We used the above method to expand  $P_n$  and calculated  $N + 1$  expansion coefficients  $a_j$ , then calculated the factorial moment of continuous order  $F_q$ . In the calculation, the program package called MATHEMATICA is adopted to ensure numerical accuracy. The results are shown in Fig. 1 with a solid line. It can be seen that  $F(q)$  has a constant value of 1. In an experiment, the access to an accurate value of  $P_n$  would require measurement of an infinite number of events. However, in a real experiment,



**Fig. 1**

The results obtained by Hwa's method and by the maximum likelihood method for Poisson distribution ( $\langle n \rangle = 1$ ).

Solid line: the results of the analytic  $P_n$ ; long-short line, dotted line, dashed line: the results for  $10^5$ ,  $10^6$ ,  $10^7$  MC events; dot-dashed line: the calculated results using the maximum likelihood method for  $10^5$  MC events sample.

one can only measure a finite number of events. So each  $P_n$  is measured with a statistical error. In order to see the effect of the statistical fluctuation, we performed a Monte Carlo simulation according to the Poisson distribution. Three samples are generated. Then we counted up the  $P_n$  and calculated  $F(q)$  according to the Hwa's method. The results are shown in Fig. 1. It can be seen from Fig. 1 that the  $F(q)$ 's deviated from 1 obviously and the deviations cannot be reduced only by increasing the event number. It is because the expansion (7) demands penetrating all the experimental points  $P_n$  precisely. So large statistical noises are included in the expansion coefficients  $a_j$ , i.e., a set of very large  $a_j$  values with alternating signs may appear and cause great instability for the results of  $F(q)$ . The  $a_j$ 's of the  $10^5$  event sample are listed in Table 1.

#### 4. IMPROVEMENT OF HWA'S METHOD

Because the negative binomial distribution can fit the experimental results well, we can choose a smaller number of  $a_j$  ( $j = 0, 1, \dots, J, J \leq N$ ) to fit the experimental data

$$\tilde{P}_n = \sum_{j=0}^J a_j P_n^{\text{NB}}(k_j, x_j), \quad (16)$$

where

$$x_j = x(1 + \Delta_j); \quad k_j = k(1 + \Delta_j);$$

$$x = \langle n \rangle = \sum_{n=0}^N n P_n;$$

$$k = (F_2 - 1)^{-1}, \quad F_2 = \langle n(n-1) \rangle / x^2;$$

$$\Delta_j = \Delta \left( -\frac{1}{2} + \frac{j}{J} \right), \quad j = 0, 1, \dots, J. \quad (17)$$

**Table 1**  
The calculated results of a  $10^5$  events sample for the Poisson distribution with  $\langle n \rangle = 1$ .

$n$	$N_n$	$N_n(\text{MC})$	$\tilde{N}_n$	$j$	$a_j$	$\tilde{a}_j$
0	36787.9	36571	36659.0	0	$-6.0310967075 \times 10^5$	-0.0436
1	36787.9	37130	36887.5	1	$5.1107712192 \times 10^6$	1.0867
2	18394.0	18274	18456.9	2	$-1.8943023020 \times 10^7$	-0.0431
3	6131.3	6120	6125.1	3	$4.0111196201 \times 10^7$	
4	1532.8	1558	1516.5	4	$-5.3070399844 \times 10^7$	
5	306.6	293	298.7	5	$4.4927260786 \times 10^7$	
6	51.1	41	48.7	6	$-2.3765120027 \times 10^7$	
7	7.3	12	6.8	7	$7.1816970312 \times 10^6$	
8	0.9	1	0.8	8	$-9.4927165712 \times 10^5$	

Notes:  $N_n$ : the event number distribution calculated according to Eq. (14);  $N_n(\text{MC})$ : the event number distribution for the Monte Carlo sample;  $\tilde{N}_n$ : the fit values of the MC sample using the maximum likelihood method;  $a_j$ : the coefficient obtained by using Hwa's method to expand  $N_n(\text{MC})$ ;  $\tilde{a}_j$ : the coefficient obtained by using the maximum likelihood method to fit  $N_n(\text{MC})$ .

$\Delta$  can be set to 0.5 [3] and the  $\Delta_j$ 's range from  $-\Delta/2$  to  $\Delta/2$  in equal steps. The fit was performed for  $N + 1$  points  $P_n$  and using the maximum likelihood method. The  $a_j$ 's are chosen so that the following likelihood function  $L$  reached its maximum,

$$L = \frac{N_{\text{ev}}!}{\prod_{n=0}^N N_n!} \prod_{n=0}^N (N_{\text{ev}} \tilde{P}_n)^{N_n}, \tag{18}$$

where  $N_{\text{ev}}$  is the total number of events, and  $N_n$  is the number of events with multiplicity  $n$ . From Eq. (16), one obtains

$$\sum_{j=0}^J a_j = 1. \tag{19}$$

which should be fulfilled in the whole process of finding the proper values of  $a_j$ . At first, we set all  $a_j$ 's to  $1/J$ . Then, we change each  $a_j$  in turn by step  $h = 1/2J$  to acquire maximum  $L$  value. When the best values of  $a_j$  are obtained at step  $h$ , we reduce the step  $h$  and find the better values of  $a_j$  at the new step. The change direction could be positive or negative. If a change is made for one  $a_j$ , an extra factor  $1/(1 + h)$  or  $1/(1 - h)$  should be multiplied to each  $a_j$  in order to fulfill Eq. (19). The last step used in our calculation is 0.0001. Usually, a suitable number of  $a_j$  is chosen according to the experimental multiplicity distribution when the fit was performed. The dot-dashed line in Fig. 1 is the result of  $J = 2$  for the Monte Carlo sample which includes  $10^5$  events. The fit values ( $\tilde{N}_n$ ) of the event number distribution are listed in Table 1. It can be seen from Fig. 1 that the maximum likelihood method can well repress the effect of the measurement error of  $P_n$ . We also calculated factorial moment of continuous order for the distribution [3]

$$P_n^{(2)} = (n + 1)^{0.5} e^{-n} / Z \tag{20}$$

where  $Z$  is a normalization factor. Unlike the Poissonian distribution, there exist not only statistical fluctuations, but also dynamic fluctuations in  $P_n^{(2)}$  distribution. At first, we generate a sample of  $10^5$  Monte Carlo events according the  $P_n^{(2)}$  distribution. Then we count up the multiplicity distribution  $P_n$

Table 2

The calculated results of a  $10^5$  Monte Carlo event sample ( $P_n^{(2)}$  distribution).

$n$	$N_n^{(2)}$	$N_n(\text{MC})$	$\tilde{N}_n$	$q$	$F(q)$
0	52016.2	52231	52206.9	-1.0	$7.7765 \times 10^{18}$
1	27061.9	26955	27047.2	-0.8	$2.1717 \times 10^{15}$
2	12193.0	12137	12103.6	-0.6	$1.5422 \times 10^{13}$
3	5179.5	5230	5124.6	-0.4	$3.1475 \times 10^{11}$
4	2130.3	2043	2107.2	-0.2	$9.1117 \times 10^9$
5	858.5	850	850.9	0.0	0.9978
6	341.1	343	339.3	0.2	$-8.5623 \times 10^7$
7	134.2	132	134.1	1.0	1.0000
8	52.3	51	52.7	1.2	$5.1967 \times 10^4$
9	20.3	14	20.6	2.0	0.7846
10	7.8	10	8.0	3.0	-18.959
11	3.0	2	3.1	4.0	-477.21
12	1.2	1	1.2	5.0	$-9.8628 \times 10^3$
13	0.4	1	0.5	6.0	$-1.9438 \times 10^5$

Notes:  $N_n^{(2)}$ : the event number distribution calculated according to Eq. (20);  $N_n(\text{MC})$ : the event number distribution for the Monte Carlo sample;  $\tilde{N}_n$ : the fit values of the Monte Carlo sample using the maximum likelihood method.

and calculate  $a_j$  and  $F(q)$  by the maximum likelihood method.  $F(q)$  are shown in Fig. 2 with solid circles. We have also applied Hwa's method to expand this Monte Carlo sample  $P_n$  and calculated the  $a_j$  and  $F(q)$ . The results are listed in Table 2. It can be seen from Table 2 that the  $F(q)$  is totally covered by statistical noise. The calculated results obtained by using Hwa's method to analytical  $P_n$  are also shown in Fig. 2 in order to compare with the results of the maximum likelihood method for the Monte Carlo sample. It can be seen that they are coincident with each other in the range  $q > 0$ , while in the range  $q < 0$ , there are deviations and the deviation is much greater for large  $N$ . The above results show that the original Hwa's method has distinct defect.

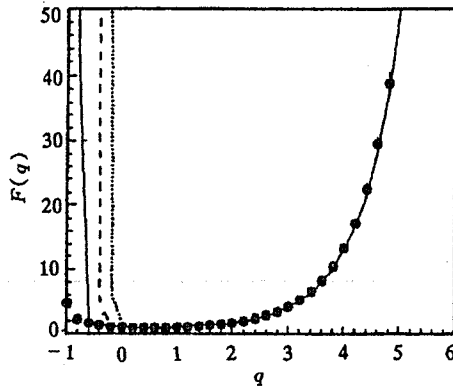


Fig. 2

The calculated results of  $F(q)$  for the  $P_n^{(2)}$  distribution.

Black circle: the calculated results using the maximum likelihood method for the MC events sample. Solid line, dashed line, dotted line: the calculated using Hwa's method for analytic  $P_n^{(2)}$  with  $N = 10, 15, \text{ and } 20$ , respectively.

### 5. EXPERIMENTAL RESULT OF THE FACTORIAL MOMENT OF CONTINUOUS ORDER $F(q)$

Using the LEBC films offered by the CERN NA27 Collaboration, we measured the pseudorapidity distribution of charged particles produced in 400 GeV/c pp collisions. 3730 non-single-diffractive events were measured. The details about the measurement are described elsewhere [4]. The fractal behavior of multiplicity production was investigated by using the method of factorial moment of continuous order as mentioned above. In order to eliminate the influence of uneven pseudorapidity distribution, the following normalized pseudorapidity  $x(\eta)$  [5] is used in the analysis

$$x(\eta) = \int_{\eta_{\min}}^{\eta} \rho(\eta') d\eta' / \int_{\eta_{\min}}^{\eta_{\max}} \rho(\eta') d\eta'. \quad (21)$$

where  $[\eta_{\min}, \eta_{\max}]$  is chosen to be  $[-2, 2]$  and  $x$  is uniformly distributed in  $[0, 1]$ .  $x$  space is divided into  $M$  bins with equal size  $\delta$ . Counting up the experimental multiplicity distribution  $P_{n, m}$  in bin  $m$ ,

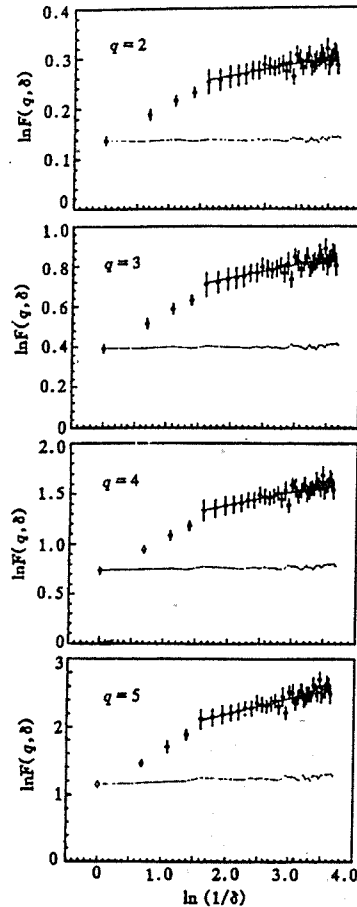


Fig. 3

The experimental results of the  $F(q)$  obtained by the maximum likelihood method.

Black points: the experimental results; solid line: the results of fitting the  $F(q, \delta)$  with Eq. (23); dotted line: the calculated results for MC events.

**Table 3**A comparison of  $D(q)$  obtained using different methods.

$q$	$D(q)(F(q))$	$D(q)(F(q \text{ order}))$	$D_{qn}^d(G_q \text{ order})$
0	1		
1	$0.9868 \pm 0.0004$		
2	$0.9764 \pm 0.0060$	$0.973 \pm 0.003$	$0.963 \pm 0.004$
3	$0.9654 \pm 0.0076$	$0.968 \pm 0.005$	$0.959 \pm 0.005$
4	$0.9535 \pm 0.0087$	$0.968 \pm 0.009$	$0.963 \pm 0.007$
5	$0.9400 \pm 0.0094$	$0.942 \pm 0.015$	$0.944 \pm 0.009$

calculation  $P_{n,m}^{NB}(x_{j,m}, k_{j,m})$  by (8), determining  $a_{j,m}$  ( $j = 0, \dots, J$ ) by the maximum likelihood method according to Eqs. (16) and (18) ( $J = 4$  for  $M \leq 4$  and  $J = 2$  for  $M > 4$ ), then  $f_m(q)$  is calculated for bin  $m$  according to Eq. (12). At last, making an average over each bin

$$F(q, \delta) = \frac{1}{M} \sum_{m=1}^M \frac{f_m(q)}{[f_m(1)]^q}, \quad (22)$$

the factorial moment  $F(q, \delta)$  of continuous order  $q$  at bin size  $\delta$  is obtained. The results are shown in Fig. 3 with solid points. In order to see the statistical contribution to  $F(q, \delta)$  we made a sample of Monte Carlo events. Comparing to the experimental data, the sample of Monte Carlo events has the same multiplicity distribution in the  $x$  space but no correlation. For event  $i$  with  $n_i$  particles, we distribute these particles randomly through  $x$  space with uniform distribution. A total of  $100N_{ev}$  events have been simulated. The calculated results are also shown in Fig. 3 with dotted lines. It can be seen that if there is only statistical fluctuation, the  $F(q, \delta)$  remain constant approximately when  $\delta$  is decreased and hence the intermittent exponents  $\phi(q)$  equal zero. This indicated that the statistical fluctuations are filtered out.

When  $\delta \rightarrow 0$ , we can get the intermittent exponent  $\phi(q)$  by fitting  $F(q, \delta)$  with the following formula (for  $M > 4$ )

$$F(q, \delta) \approx \delta^{-\phi(q)} \quad (23)$$

Then we can calculate multifractal dimension and multifractal spectrum using the relations

$$\begin{aligned} \tau(q) &= q-1-\phi(q), \quad D(q) = \tau(q) / (q-1); \\ \alpha &= d\tau(q) / dq, \quad f(\alpha) = q\alpha - \tau(q). \end{aligned} \quad (24)$$

The multifractal dimension  $D(q)$  versus  $q$  obtained from experiment is shown in Fig. 4. It can be seen from Fig. 4 that the  $D(q)$  is monotonously decreased with increasing  $q$ . This means that there is multifractal behavior in the process of multiplicity production in pp collisions. The results of the Monte Carlo samples with particles randomly distributed in pseudorapidity space are shown in Fig. 4 with a solid line. It can be seen that  $D(q) = 1$  at each  $q$  for the Monte Carlo samples approximately. It means that the statistical fluctuation is filtered out by the method of maximum likelihood.

The calculated values of  $D(q)$  at integer  $q$  are listed in Table 3 in comparison with that obtained by the ordinary scaled factorial moment and modified  $G$  moments. It can be seen that they are well consistent with each other. So the method of the continuous order of factorial moment is successful. The results of multifractal spectrum  $f(\alpha)$  are shown in Fig. 5. It is a convex curve with a maximum at  $q = 0$ ,  $f(\alpha(0)) = D(0) = 1$ . The straight line  $f(\alpha) = \alpha$  is tangent to the  $f(\alpha)$  curve at  $q = 1$ . The black point represents the results of the Monte Carlo sample, which essentially condense to a single



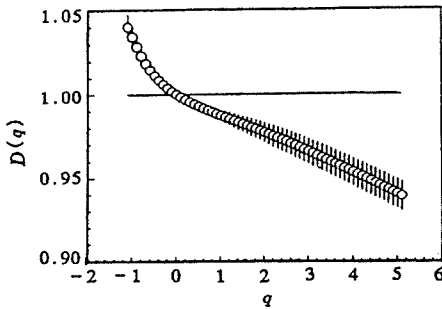


Fig. 4

Fractal dimension  $D(q)$  versus  $q$ .

Open circle: the experimental value; solid line: the result of the Monte Carlo events.

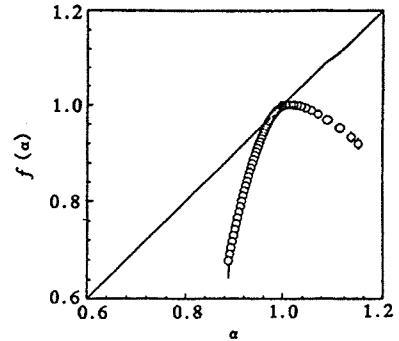


Fig. 5

Multifractal spectrum  $f(\alpha)$  versus  $\alpha$ .

Open circle: the experimental results; black circle: the results of the Monte Carlo events.

point,  $a = f(\alpha) = 1$ . The experimental multifractal spectrum is not a point showing that there is multifractal behavior in multiparticle production at 400 GeV/c pp collisions.

## 6. CONCLUSION

The method of the factorial moments  $F_q$  of continuous order suggested by Hwa has been tested. It is found that using this method to analyze the experimental data cannot obtain satisfactory results. Some improvements have been made for Hwa's method, which makes it suitable for the analysis of experimental data. The analytic results for the experimental pseudorapidity distributions of charged particles produced in 400 GeV/c pp collisions indicated that the method of the factorial moments of continuous order is feasible and correct. There is possibly multifractal behavior in the process of multiplicity production in pp collisions at 400 GeV/c.

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